

Determination of the Exact Optimum Order Statistics for Estimating the Parameters of the Exponential Distribution from Censored Samples*

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This paper presents the small sample optimum choice of the $k(\leq r_2)$ order statistics for the best linear unbiased estimate (BLUE) of the parameters μ and σ or σ alone (μ known) when the sample is Type II censored on the right. For $n=2(1)10$, $k=1(1)r_2$ and $r_2 = \{[.50n]+1\} (1)n$, the optimum ranks, the coefficients of the BLUEs have been presented in Table I.

INTRODUCTION

Assume that we are sampling from the exponential distribution

$$f(x) = \frac{1}{\sigma} e^{-(x-\mu)/\sigma} \quad x \geq \mu; \quad \sigma > 0 \quad (1.1)$$

under Type II censoring procedure with a fixed proportion of censoring on the right. The symbols μ and σ denote the location and the scale parameters respectively. From a random sample of size n let $x_{(1)} < x_{(2)} < \dots < x_{(r_2)}$ be the uncensored portion of the sample. The integer r_2 is the rank of the largest observation in the uncensored portion of the sample. We are interested in the estimation of μ and σ or σ alone, when μ is known, on the basis of k suitably chosen order statistics $x_{(n_1)}, x_{(n_2)}, \dots, x_{(n_k)}$ where n_1, n_2, \dots, n_k are the ranks satisfying the inequality $1 \leq n_1 < \dots < n_k \leq r_2$.

The small sample as well as the asymptotic ($n \rightarrow \infty$) situation in uncensored samples (i.e., all the n observations are available for estimation) has been considered by Harter (1961), Kulldorff (1963a, 1963b), Ogawa (1960), Saleh (1964), Saleh and Ali (1966a), Sarhan, Greenberg and Ogawa (1963) and Siddiqui (1963). The asymptotic ($n \rightarrow \infty$) situation concerning the optimum choice of the $k(\leq r_2)$ order statistics has been considered by Saleh (1964, 1966b).

We shall present here the small sample optimum choice of the $k(\leq r_2)$ order statistics for the best linear unbiased estimate (BLUE) of the parameters μ and σ or σ alone (μ known) when the sample is Type II censored on the right. For $n = 2(1)10$, $k = 1(1)r_2$ and $r_2 = \{[.50n] + 1\} (1)n$, the optimum ranks, the coefficients of the BLUEs have been presented in Table I.

Received November 1965. Revised July 1966.

* This paper has been completed with partial support from the National Research Council of Canada.

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TABLE 1: The Coefficients of Blue of the Scale Parameter Based on Optimum Order Statistics From Censored Samples

$n = 2$						$n = 3$					
r_2	k	n_1	n_2	K^*		r_2	k	n_1	n_2	n_3	K^*
1	1	1		1.0000		2	1	2			1.9231
		2.0000					2	.5455			
2	1	2		1.8000			2	2	3		2.9231
		.6667				3	1	.4474	.3421		
							3	3			2.4694
2	2	1	2	2.0000			2	.5455			
		.5000	.5000				2	2	3		2.9231
							3	.4474	.3421		
							3	1	2	3	3.0000
								.3333	.3333	.3333	

TABLE I Continued

$n = 4$

r_2	k	n_1	n_2	n_3	n_4	K^*
2	1	2				1.9600
		1.7143				
2	1	1	2			2.0000
		.5000	1.5000			
3	1	3				2.7705
		.9231				
2	2	2	3			2.9600
		.4595	.6757			
3	1	1	2	3		3.0000
		.3333	.3333	.6667		
4	1	4				3.0488
		.4800				
2	2	3	4			3.7705
		.4130	.2652			
3	3	2	3	4		3.9600
		.3434	.2525	.2525		
4	1	1	2	3	4	4.0000
		.2500	.2500	.2500	.2500	

$n = 5$

r_2	k	n_1	n_2	n_3	n_4	n_5	K^*
3	1	3					2.8726
		1.2766					
1	1	2	3				2.9756
		.4672	1.0082				
3	1	1	2	3			3.0000
		.3333	.3333	1.0000			
4	1	4					3.5524
		.7792					
2	2	2	4				3.8987
		.5342	.5919				
3	2	2	3	4			3.9756
		.3497	.2515	.5031			
5	1	5					3.5622
		.4380					
2	2	3	5				4.6726
		.5280	.2568				
3	2	2	4	5			4.8987
		.4251	.2669	.2041			
4	2	2	3	4	5		4.9757
		.2794	.2010	.2010	.2010		
5	1	1	2	3	4	5	5.0000
		.2000	.2000	.2000	.2000	.2000	

* The original calculations were carried with precision to seven decimal places.

TABLE I *Continued*

$n = 6$

r_2	k	n_1	n_2	n_3	n_4	n_5	n_6	K^*
3	1	3 1.6216						2.4190
	2	2 .4725	3 1.3407					2.9836
	3	1 .3333	2 .3333	3 1.3333				3.0000
4	1	4 1.0526						4.2787
	2	2 .5198	4 .8520					3.9436
	3	2 .3539	3 .2510	4 .7531				3.9836
	4	1 .2500	2 .2500	3 .2500	4 .7500			4.9836
5	1	5 .6897						4.2787
	2	3 .5010	5 .4766					5.5388
	3	2 .4146	4 .2751	5 .4046				4.9436
	4	2 .2829	3 .2007	4 .2007	5 .4013			4.9836
	5	1 .2000	2 .2000	3 .2000	4 .2000	5 .4000		5.0000
6	1	5 .6897						4.2787
	2	4 .4939	6 .2167					5.5388
	3	3 .4152	5 .2238	6 .1712				5.8421
	4	2 .3449	4 .2288	5 .1682	6 .1682			5.9436
	5	2 .2356	3 .1671	4 .1671	5 .1671	6 .1671		5.9836
	6	1 .1667	2 .1667	3 .1667	4 .1667	5 .1667	6 .1667	6.0000

TABLE I *Continued*

$n = 8$

r_2	k	n_1	n_2	n_3	n_4	n_5	n_6	n_7	n_8	K^*
4	1	4								3.8784
		1.5760								
	2	2	4							3.9748
		.5092	1.3610							
3	2	3	4							3.9912
		.3592	.2506	1.2528						
4	1	2	3	4						4.0000
		.2500	.2500	.2500	1.2500					
5	1	5								4.7043
		1.1305								
	2	3	5							4.9345
		.4903	.8897							
	3	2	4	5						4.9748
	.4068	.2834	.8041							
4	2	3	4	5						4.9912
		.2872	.2004	.2004	.8014					
5	1	2	3	4	5					5.0000
		.2000	.2000	.2000	.2000	.8000				
6	1	6								5.3463
		.8211								
	2	4	6							5.8384
		.4714	.5755							
	3	2	4	6						5.9348
		.3410	.3454	.5662						
4	2	4	5	6						5.9748
		.3387	.2360	.1674	.5021					
5	2	3	4	5	6					5.9912
		.2393	.1669	.1669	.1669	.5007				
6	1	2	3	4	5	6				6.0000
		.1667	.1667	.1667	.1667	.1667	.5000			
7	1	7								5.5952
		.5821								
	2	4	7							6.6489
		.5347	.3846							
	3	3	5	7						6.8489
		.3528	.3037	.3365						
	4	2	4	6	7					6.9348
	.2918	.2956	.1961	.2884						
5	2	4	5	6	7					6.9748
		.2902	.2021	.1434	.1434	.2867				
6	2	3	4	5	6	7				6.9912
		.2051	.1430	.1430	.1430	.1430	.2861			
7	1	2	3	4	5	6	7			7.0000
		.1429	.1429	.1429	.1429	.1429	.1429	.2857		
8	1	7								5.5952
		.5821								
	2	5	8							7.1737
		.5536	.1878							
3	4	7	8							7.6489
		.4648	.2036	.1307						
4	3	5	7	8						7.8576
		.3079	.2650	.1664	.1273					

TABLE I *Continued*

$n = 8$ Continued —

r_2	k	n_1	n_2	n_3	n_4	n_5	n_6	n_7	n_8	K^*
8	5	2	4	6	7	8				7.9348
		.2551	.2583	.1714	.1260	.1260				
	6	2	4	5	6	7	8			7.9748
		.2538	.1768	.1254	.1254	.1254	.1254			
	7	2	3	4	5	6	7	8		7.9912
		.1794	.1251	.1251	.1251	.1251	.1251	.1251		
	8	1	2	3	4	5	6	7	8	8.0000
		.1250	.1250	.1250	.1250	.1250	.1250	.1250	.1250	

$n = 9$

r_2	k	n_1	n_2	n_3	n_4	n_5	n_6	n_7	n_8	n_9	K^*
5	1	5									4.7864
		1.3411									
	2	3	5								4.9522
		.4894	1.0924								
	3	2	4	5							4.9813
		.4051	.2858	1.0037							
	4	2	3	4	5						4.9931
		.2887	.2003	.2003	1.0014						
	5	1	2	3	4	5					5.0000
		.2000	.2000	.2000	.2000	1.0000					
6	1	6									5.5486
		1.0044									
	2	3	6								5.8876
		.5265	.8040								
	3	2	4	6							5.9569
		.3387	.3413	.7370							
	4	2	4	5	6						5.9313
		.3374	.2380	.1672	.6637						
	5	2	3	4	5	6					5.9931
		.2405	.1669	.1669	.1669	.6674					
	6	1	2	3	4	5	6				6.0000
		.1667	.1667	.1667	.1667	.1667	.6667				
7	1	7									6.0951
		.7525									
	2	4	7								6.7818
		.5157	.5407								
	3	3	5	7							6.9122
		.3506	.2966	.4861							
	4	2	4	6	7						6.9569
		.2900	.2923	.1998	.4312						
	5	2	4	5	6	7					6.9813
		.2890	.2039	.1432	.1432	.4297					
	6	2	3	4	5	6	7				6.9931
		.2061	.1430	.1430	.1430	.1430	.4290				
	7	1	2	3	4	5	6	7			7.0000
		.1429	.1429	.1429	.1429	.1429	.1429	.4286			

TABLE I *Continued*

$n = 10$

r_2	k	n_1	n_2	n_3	n_4	n_5	n_6	n_7	n_8	n_9	n_{10}	K^*	
5	1	5										4.8382	
	2	1.5489											
		3	5										4.9634
	3	.4892	1.2942										
		2	4	5									4.9856
4	.4039	.2876	1.2035										
	2	3	4	5								4.9945	
5	.2898	.2002	.2002	1.2013									
	1	2	3	4	5							5.0000	
6	.2000	.2000	.2000	.2000	1.2000								
	1	6										5.6683	
	2	1.1825											
		3	6										5.9192
	3	.5193	.9761										
		2	4	6									5.9692
4	.3374	.3390	.9063										
	2	4	5	6								5.9856	
5	.3365	.2395	.1671	.8353									
	2	3	4	5	6							5.9945	
6	.2415	.1668	.1668	.1668	.8341								
	1	2	3	4	5	6						6.0000	
7	.1667	.1667	.1667	.1667	.1667	.8333							
	1	7										6.3630	
	2	.9127											
		4	7										6.8487
	3	.5068	.6912										
		3	5	7									6.9391
	4	.3500	.2930	.6327									
2		4	6	7								6.9692	
5	.2890	.2904	.2023	.5740									
	2	4	5	6	7							6.9856	
6	.2883	.2052	.1432	.1432	.5726								
	2	3	4	5	6	7						6.9945	
7	.2070	.1430	.1430	.1430	.1430	.5719							
	1	2	3	4	5	6	7					7.0000	
8	.1429	.1429	.1429	.1429	.1429	.1429	.5714						
	1	8										6.8118	
	2	.6998											
		5	8										7.7108
	3	.4963	.4756										
		3	6	8									7.8792
	4	.3901	.3069	.4264									
		3	5	7	8								7.9391
5	.3059	.2561	.1751	.3779									
	2	4	6	7	8							7.9692	
6	.2527	.2540	.1769	.1255	.3764								
	2	4	5	6	7	8						7.9856	
7	.2522	.1795	.1252	.1252	.1252	.3757							
	2	3	4	5	6	7	8					7.9945	
8	.1811	.1251	.1251	.1251	.1251	.1251	.3753						
	1	2	3	4	5	6	7	8				8.0000	
		.1250	.1250	.1250	.1250	.1250	.1250	.1250	.3750				

TABLE I *Continued*

$n = 10$ Continued

r_2	k	n_1	n_2	n_3	n_4	n_5	n_6	n_7	n_8	n_9	n_{10}	K^*
9	1	8										6.8118
		.6998										
	2	6	9									8.4388
		.4913	.3030									
	3	4	7	9								8.7717
		.3957	.2765	.2631								
	4	3	6	8	9							8.8792
		.3462	.2723	.1532	.2252							
	5	3	5	7	8	9						8.9391
	.2717	.2275	.1555	.1119	.2237							
6	2	5	6	7	8	9					8.9692	
	.2245	.2256	.1572	.1115	.1115	.2230						
7	2	4	5	6	7	8	9				8.9856	
	.2241	.1595	.1113	.1113	.1113	.1113	.2226					
8	2	3	4	5	6	7	8	9			8.9945	
	.1609	.1112	.1112	.1112	.1112	.1112	.1112	.2224				
9	1	2	3	4	5	6	7	8	9		9.0000	
	.1111	.1111	.1111	.1111	.1111	.1111	.1111	.1111	.2222			
10	1	8										6.8118
		.6998										
	2	7	10									8.8324
		.5050	.1525									
	3	5	8	10								9.5108
		.4023	.2594	.1262								
	4	4	7	9	10							9.7717
		.3552	.2482	.1338	.1023							
	5	3	6	8	9	10						9.8792
		.3112	.2447	.1377	.1012	.1012						
6	3	5	7	8	9	10					9.9391	
	.2443	.2046	.1399	.1006	.1006	.1006						
7	2	4	6	7	8	9	10				9.9692	
	.2020	.2030	.1414	.1003	.1003	.1003	.1003					
8	2	4	5	6	7	8	9	10			9.9856	
	.2017	.1435	.1001	.1001	.1001	.1001	.1001	.1001				
9	2	3	4	5	6	7	8	9	10		9.9945	
	.1448	.1001	.1001	.1001	.1001	.1001	.1001	.1001	.1001			
10	1	2	3	4	5	6	7	8	9	10	10.0000	
	.1000	.1000	.1000	.1000	.1000	.1000	.1000	.1000	.1000	.1000		

OPTIMUM RANKS OF ORDER STATISTICS FOR THE ESTIMATION OF σ (μ KNOWN)

For $k(\leq r_2)$ let $x_{(n_1)}, x_{(n_2)}, \dots, x_{(n_k)}$ be the k order statistics of fixed ranks n_1, n_2, \dots, n_k satisfying the inequality $1 \leq n_1 < \dots < n_k \leq r_2$. Then the BLUE of σ due to Kuldorff (1963) is, $\hat{\sigma} = b_0\mu + \sum_{i=1}^k b_i x_{(n_i)}$ where

$$b_i = (\delta_{1,i}/\delta_{2,i} - \delta_{1,i+1}/\delta_{2,i+1})/K \text{ for } i = 0, 1, \dots, k$$

and

$$\delta_{r,i} = \sum_{j=n_{i-1}}^{n_i-1} (n-j)^{-r} \quad (r = 1, 2; i = 1, 2, \dots, k);$$

$$K = \sum_{i=1}^k \delta_{1,i}^2 / \delta_{2,i}$$

The variance of $\hat{\sigma}$ is $V(\hat{\sigma}) = \sigma^2/K$.

In order to determine the optimum ranks of the order statistics, we minimize

the variance of σ or, equivalently, maximize K subject to the restriction $1 \leq n_1 < \dots < n_k \leq r_2$. The maximization of K subject to the restriction $1 \leq n_1 < \dots < n_k \leq n$ has been considered by Kulldorff (1963b). Let this maximum be attained at $(n_1^0 < n_2^0 < \dots < n_k^0)$. For the maximum of K we have two situations (i) the proportion of censoring is such that $r_2 \geq n_k^0$; (ii) the proportion of censoring is such that $r_2 < n_k^0$. In the first case, the optimum ranks are n_1^0, \dots, n_k^0 . In the second case, we have solved the problem numerically on an IBM 7040 for $n = 2(1)10, r_2 = \{[.50n] + 1\}(1)n$ and $k = 1(1)r_2$. The optimum ranks, the coefficients of the BLUE and the maximum value of K are presented in Table I. If we use the optimum ranks, the relative efficiency (RE) of this BLUE of σ compared to the BLUE based on all the observations in the censored sample (Sarhan and Greenberg, 1957) is K^*/r_2 , where K^* is the maximum of K , subject to the restriction $1 \leq n_1 < \dots < n_k \leq r_2$.

OPTIMUM RANKS OF ORDER STATISTICS FOR THE ESTIMATION OF μ AND σ

For $k(\leq r_2)$ let $x_{(n_1)}, \dots, x_{(n_k)}$ be k order statistics of fixed ranks n_1, \dots, n_k satisfying the inequality $1 \leq n_1 < \dots < n_k \leq r_2$. Then the BLUE's of σ and μ due to Kulldorff (1963b) are:

$$\hat{\sigma} = \sum_{i=1}^k b_i x_{(n_i)}$$

$$\hat{\mu} = x_{(n_1)} - \hat{\sigma} \delta_{11}$$

where

$$b_i = -\delta_{12}/L \delta_{22} \text{ for } i = 1$$

$$= (\delta_{1,i}/\delta_{2,i} - \delta_{1,i+1}/\delta_{2,i+1})/L \text{ for } i = 2, \dots, k$$

and

$$L = \sum_{i=2}^k \delta_{1,i}^2/\delta_{2,i}$$

The variances and covariance of the BLUE's are

$$V(\hat{\sigma}) = \sigma^2/L$$

$$V(\hat{\mu}) = \sigma^2(\delta_{21} + \delta_{11}^2/L)$$

$$\text{Cov}(\hat{\mu}, \hat{\sigma}) = -\sigma^2 \delta_{11}/L$$

The generalized variance of the estimates $\hat{\mu}$ and $\hat{\sigma}$ is

$$\Lambda(\hat{\mu}, \hat{\sigma}) = V(\hat{\mu})V(\hat{\sigma}) - \text{Cov}^2(\hat{\mu}, \hat{\sigma})$$

$$= \sigma^4 \delta_{21}/L$$

In order to determine the optimum ranks of the order statistics, we minimize the generalized variance $\sigma^4 \delta_{21}/L$ with respect to n_1, n_2, \dots, n_k or equivalently maximize $L(\delta_{21})^{-1}$ subject to the restriction $1 \leq n_1 < \dots < n_k \leq r_2$. By theorem 7 (Kulldorff, 1963b), it is clear that $L(\delta_{21})^{-1}$ as a function n_1 decreases;

therefore we choose $n_1 = 1$. Let us now turn to the second step of maximizing $L(\delta_{21})^{-1}$ with respect to n_2, \dots, n_k while keeping $n_1 = 1$ fixed. If we substitute $n - 1$ for n , $k - 1$ for k and $n_{i+1} - 1$ for n'_i ($i = 1, 2, \dots, k - 1$) in the expression for K in Section II we obtain

$$K' = \sum_{i=1}^{k-1} \frac{\delta'_{1i}}{\delta'_{2i}}$$

where

$$\delta'_{r,i} = \sum_{j=n'_{i-1}}^{n'_{i+1}-1} (n - 1 - j)^{-r} \quad (r = 1, 2; i = 1, 2, \dots, k - 1)$$

Therefore $L = n^2 K'$, where K' is a function of $n'_1, n'_2, \dots, n'_{k-1}$ satisfying the inequality $1 \leq n'_1 < \dots < n'_{k-1} \leq (r_2 - 1)$. The remaining problem of maximizing K' with respect to $n'_1 \dots n'_{k-1}$ subject to $1 \leq n'_1 < \dots < n'_{k-1} \leq (r_2 - 1)$ has been discussed in Section II and we encounter two situations (i) $(r_2 - 1) \geq n^*_{k-1}$ (ii) $(r_2 - 1) < n^*_{k-1}$ where n^*_{k-1} is the largest rank in selecting $k - 1$ order statistics from an uncensored sample of size $n - 1$.

In the first case, the optimum ranks coincide with the ranks when the sample is not censored. In the second case, we maximize K' with respect to $n'_1 \dots n'_{k-1}$ numerically subject to the restriction $1 \leq n'_1 < \dots < n'_{k-1} \leq (r_2 - 1)$. Let these ranks be $n^{**}_1, \dots, n^{**}_{k-1}$. Then the optimum ranks of the k order statistics for the estimation of μ and σ simultaneously are

$$1, (n^{**}_1 + 1), (n^{**}_2 + 1), \dots, (n^{**}_{k-1} + 1)$$

and the estimates are

$$\begin{aligned} \hat{\mu} &= x_{(1)} - \hat{\sigma}/n \\ \hat{\sigma} &= -\left(\sum_{i=1}^{k-1} b_i^{**}\right)x_{(1)} + \sum_{i=1}^{k-1} b_i^{**}x_{(n_i^{**}+1)} \end{aligned}$$

where the coefficients b_i^{**} ($i = 1, \dots, k - 1$) are based on the Ranks 1 and $(n_i^{**} + 1)$ ($i = 1, 2 \dots k - 1$). If we use these ranks, the joint efficiency (JE) of the estimates and the relative efficiencies (RE) of the estimates of μ and σ are

$$\begin{aligned} \text{JE}(\hat{\mu}, \hat{\sigma}) &= \frac{K^*}{(r_2 - 1)} \\ \text{RE}(\hat{\sigma}) &= \frac{K^*}{(r_2 - 1)} \\ \text{RE}(\hat{\mu}) &= \frac{r_2 K^*}{(r_2 - 1)(K^* + 1)} \end{aligned}$$

where K^* is the maximum of K' .

We present two examples for the estimation of σ when μ is known and that of μ and σ when both are unknown.

Example 1: (Estimation of σ when μ is known). Assume $n = 9, r_2 = 7$ and $k = 4$. From Table I, we obtain the optimum ranks as 2, 4, 6 and 7. The BLUE of σ based on the corresponding order statistics is given by

$$\hat{\sigma} = -1.2133u + .2900x_{(2)} + .2923x_{(4)} + .1998x_{(6)} + .4312x_{(7)} .$$

The coefficient of μ is the negative of the sum of the coefficients of the order statistics in the above estimate. The Relative efficiency of the estimate is given by

$$RE(\hat{\sigma}) = \frac{6.9569}{7.0000} = 99.38\%$$

Example 2: (Estimation of μ and σ). Assume $n = 9$, $r_2 = 7$ and $k = 5$. From the reasonings of Section III we first select the first order statistics $x_{(1)}$. Then use Table I for $n = 8$, $r_2 = 6$ and $K = 4$ to obtain the optimum ranks as 2, 4, 5 and 6. Thus, the ranks of the relevant optimum order statistics to be used for estimation of μ and σ are 1, 3, 5, 6 and 7. The BLUEs are given by

$$\hat{\mu} = x_{(1)} - \hat{\sigma}/n$$

$$\hat{\sigma} = -1.2442x_{(1)} + .3387x_{(3)} + .2360x_{(5)} + .1674x_{(6)} + .5021x_{(7)}$$

Note that the coefficients of the estimate of σ are available in Table I for $n = 8$, $r_2 = 6$ and $K = 4$. The coefficient of $x_{(1)}$ is the negative sum of the coefficients of the remaining order statistics used in the estimate of σ .

The joint efficiency and relative efficiencies of the estimates are

$$JE(\hat{\mu}, \hat{\sigma}) = \frac{5.9746}{6} = 99.42\%$$

$$RE(\hat{\sigma}) = \frac{5.9746}{6} = 99.42\%$$

$$RE(\hat{\mu}) = \frac{7 \times 5.9746}{6 \times 6.9746} = 99.70\%$$

ACKNOWLEDGEMENT

The author wishes to thank Mr. F. Zabransky of the Computing Centre U. W. O. for programming of the results in Table I.

NOTE: While reviewing this paper the author has extensively computed the Tables of optimum ranks, coefficients of linear estimates and the maximum value of k for all possible right censoring. The results have been deposited with the Mathematics of Computation as unpublished Mathematical tables under the title "Tables for estimation of exponential distribution by linear combination of the optimal subset of order statistics by F. Zabransky, M. Sibuya and A. K. Md. Ehsanes Saleh."

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